

DO ATLANTA RESIDENTS VALUE MARTA? SELECTING AN AUTOREGRESSIVE MODEL TO RECOVER WILLINGNESS-TO-PAY

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HEDONIC MODELS

Understanding what homeowners are willing to pay to live close to public transit infrastructure is important for policy and forecasting. This is usually estimated with a hedonic linear regression model,

$$\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon} \quad (1)$$

with y_i the price of a home and \mathbf{x}_i its attributes, and $\boldsymbol{\beta}$ the marginal willingness-to-pay (MWTP) for the attributes. **But this model may be inadequate** under the following conditions:

Spatial Dependence The price of a home is relative to the prices of homes nearby. This creates a missing variable bias on $\hat{\boldsymbol{\beta}}$. This is remedied with a spatial autoregressive lag (SAR) model,

$$\mathbf{y} = \rho W\mathbf{y} + X\boldsymbol{\beta} + \boldsymbol{\epsilon} \quad (2)$$

Spatial Correlation Homes near each other have similar characteristics, and thus correlated errors. This invalidates OLS estimates of σ . This can be remedied with the spatial error model (SEM),

$$\mathbf{y} = X\boldsymbol{\beta} + \mathbf{u}, \mathbf{u} = \lambda W\mathbf{u} + \boldsymbol{\epsilon} \quad (3)$$

Spatial Endogeneity Unobservable neighborhood attributes are important to price. This can cause both a missing variable bias and correlated errors. This is remedied with the spatial Durbin model (SDM)

$$\mathbf{y} = \rho W\mathbf{y} + X\boldsymbol{\beta} + WX\boldsymbol{\gamma} + \boldsymbol{\epsilon} \quad (4)$$

In these models, the matrix W defines the spatial relationship between observations in the dataset. If i and j are neighbors, $[W]_{ij} \neq 0$.

Identifying the appropriate model for a given situation is an unresolved empirical problem.

EMPIRICAL APPLICATION

4,812 Targeted marketing records of households within 5 miles of a Metropolitan Atlanta Rapid Transit Authority (MARTA) rail station. Variables:

- (log) Market value of home — dependent variable
- (log) Euclidean distance to MARTA station
- (log) Property acreage
- (log) Home square footage
- (log) Age of home
- (log) Household income
- (log) Euclidean distance to freeway entrance.
- Property type (condo/single)
- Ethnicity

We estimate the models by maximum likelihood with the `spdep` package for R [3].

MODEL SELECTION

MISSPECIFICATION CONSEQUENCES

True DGP	Estimated Model			
	OLS	SAR	SEM	SDM
OLS: $\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon}$	-	inefficient	inefficient	inefficient
SAR: $\mathbf{y} = \rho W\mathbf{y} + X\boldsymbol{\beta} + \boldsymbol{\epsilon}$	$\hat{\boldsymbol{\beta}}$ biased	-	$\hat{\boldsymbol{\beta}}$ biased	inefficient
SEM: $\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\epsilon}, \boldsymbol{\epsilon} = \lambda W\boldsymbol{\epsilon} + \mathbf{u}$	$\hat{\sigma}^2$ invalid	$\hat{\sigma}^2$ invalid	-	inefficient
SDM: $\mathbf{y} = \rho W\mathbf{y} + X\boldsymbol{\beta} + WX\boldsymbol{\gamma} + \mathbf{u}$	$\hat{\boldsymbol{\beta}}$ biased	$\hat{\sigma}^2$ invalid	$\hat{\boldsymbol{\beta}}$ biased	-

CLASSICAL FRAMEWORK

Spatial models can be expensive to estimate, can OLS residuals show need?

$$H_0 : \rho, \lambda, \gamma = 0; \quad \text{No spatial effects.}$$

$$H_a : \rho \text{ or } \lambda \text{ or } \gamma \neq 0; \quad \text{Spatial effects.}$$

Test: Lagrange multipliers [1, 2]

$$LM_\rho = \frac{(\hat{\boldsymbol{\epsilon}}' W \mathbf{y} / \hat{\sigma}^2)^2}{nJ}$$

$$LM_\lambda = \frac{(\hat{\boldsymbol{\epsilon}}' W \hat{\boldsymbol{\epsilon}} / \hat{\sigma}^2)^2}{T}$$

with

$$J = \frac{1}{n\hat{\sigma}^2} [(WX\hat{\boldsymbol{\beta}})'(I - X(X'X)^{-1}X')(WX\hat{\boldsymbol{\beta}}) + T\hat{\sigma}^2]$$

and T the trace of the matrix $W'W + W^2$

GENERAL FRAMEWORK

The SDM is a general model that subsumes the other three.

$$H_0 : \rho, \lambda, \gamma \neq 0; \quad \text{Spatial effects.}$$

$$H_a : \rho \text{ or } \lambda \text{ or } \gamma = 0; \quad \text{No spatial effects.}$$

Test: likelihood ratio to estimate posterior probabilities $SAR + SEM = SDM : \pi_a + \pi_b = 1$ [4, 6]

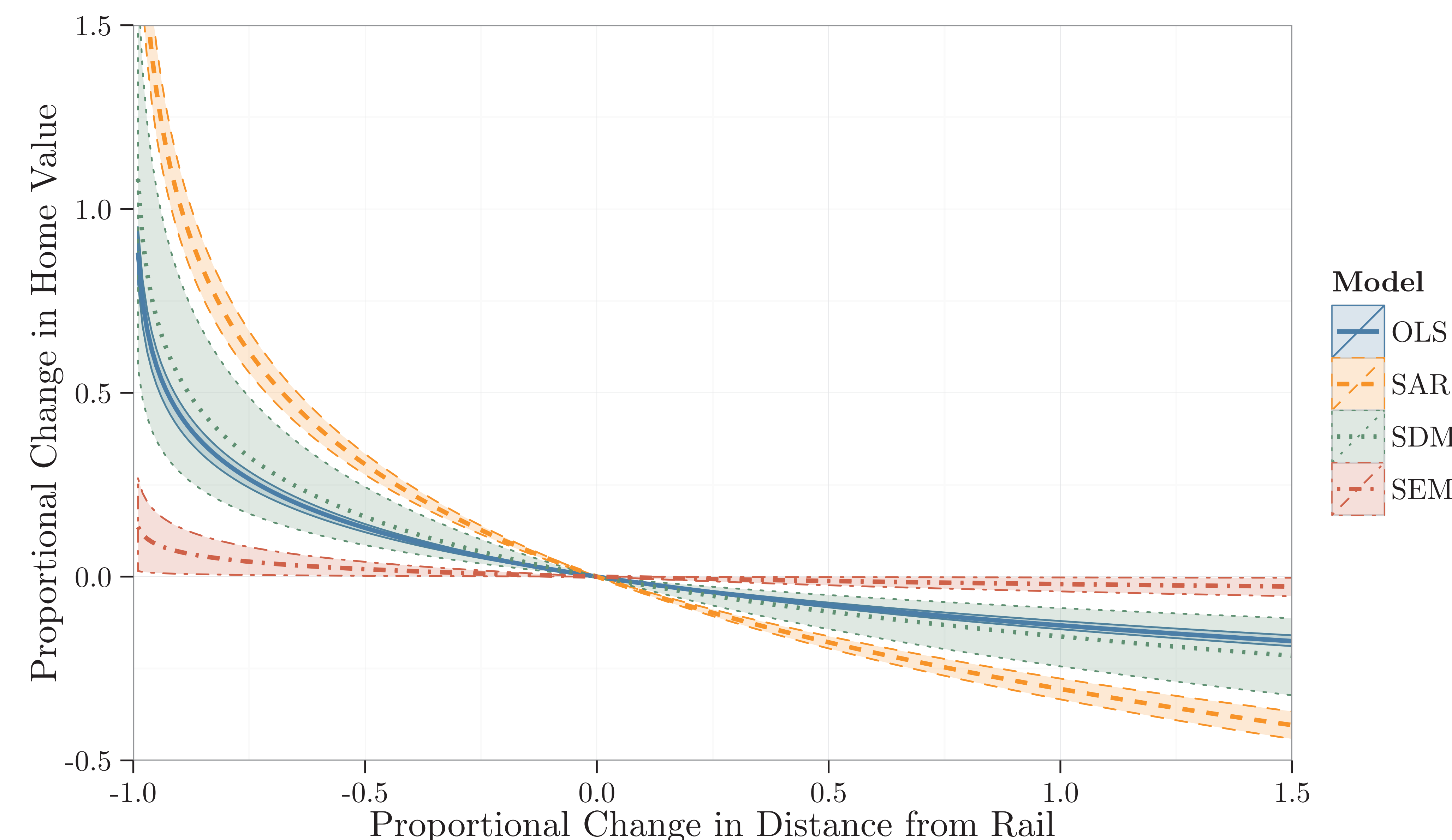
$$\mathbf{y}_c = \pi_a \mathbf{y}_a + \pi_b \mathbf{y}_b$$

$$\mathbf{y}_c = \pi_a ((I - \rho W)^{-1} (X\boldsymbol{\beta} + \boldsymbol{\epsilon})) + \pi_b (X\boldsymbol{\beta} + (I - \rho W)^{-1} \boldsymbol{\epsilon})$$

$$(I - \rho W)\mathbf{y}_c = X(\pi_a \boldsymbol{\beta}) + (I - \rho W)(X\pi_b \boldsymbol{\beta}) + (\pi_a + \pi_b)\boldsymbol{\epsilon}$$

$$\mathbf{y}_c = \rho W\mathbf{y}_c + (\pi_a + \pi_b)X\boldsymbol{\beta} + WX(-\rho\pi_b \boldsymbol{\beta}) + \boldsymbol{\epsilon}$$

ESTIMATED MWTP WITH 95% CONFIDENCE INTERVALS



OTHER STUDIES

Statistic	Osland [8]	Löchl [7]	Ibeas [5]	This Study
$\ln(\mathcal{L}_{SDM})$	96.9	4192.6	-2.1	-272.7
$\ln(\mathcal{L}_{SEM})$	80.9	4118.0	-34.8	-338.1
$\ln(\mathcal{L}_{SAR})$	65.8		-75.0	-838.8
$\ln(\mathcal{L}_{OLS})$	52.0	3183.3	-111.9	-2006.7
LR statistic	31.9*	149.2*	65.5*	130.9*
Classical General	SEM [†] SDM	SEM SDM	Unknown [‡] SDM	SEM SDM

* Reject null hypothesis with $p < 0.01$.

[†] Selected SDM after likelihood ratio test.

[‡] Did not report LM statistics, but selected SEM.

CONCLUSIONS

The econometric treatment of **spatial effects matters for policy interpretations**; the SDM estimates of MWTP are substantially less certain than the OLS estimates in this case.

Existing **selection frameworks lead to different models**; the classical framework selected the SEM which in this case appears inconsistent.

The general framework **minimizes the consequences of Type I error** error; hypothesis testing is built to minimize the risk of Type I error.

The **SDM is general and conservative**:

- It is a linear combination of the SAR and SEM models.
- It is the outcome of a spatial fixed effects process.
- It is the outcome of a spatially endogenous variables process.

The SDM may be inefficient; if researchers wish for a more efficient model they can test restrictions to the SDM.

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